

Minimax lower bound for 1-bit Matrix Completion

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Main References

- ▶ John Duchi. “Lecture notes for statistics 311/electrical engineering 377”. In: *URL: [https://stanford.edu/class/stats311/Lectures/full notes.pdf](https://stanford.edu/class/stats311/Lectures/full%20notes.pdf). Last visited on 2 [2016]*, p. 23
- ▶ Jonathan Scarlett et al. “An introductory guide to Fano’s inequality with applications in statistical estimation”. In: *arXiv preprint arXiv:1901.00555* [2019]
- ▶ Alexandre B Tsybakov. *Introduction to Nonparametric Estimation*. Springer series in statistics. Dordrecht: Springer, 2009. DOI: 10.1007/b13794. URL: <https://cds.cern.ch/record/1315296>
- ▶ Mark A Davenport et al. “1-bit matrix completion”. In: *Information and Inference: A Journal of the IMA* 3.3 [2014], pp. 189–223

Recap: General Setting

- ▶ From a distribution family $\mathcal{P} = \mathcal{N}_d = \{N(\theta, I_d) | \theta \in \mathbb{R}^d\}$, God chooses a distribution $P \in \mathcal{P}$.
- ▶ A set of N (i.i.d) samples X_1^N are drawn from P , denoted as \mathbf{X} .
- ▶ Task: estimating $\theta(P)$ from given samples.
- ▶ Quality of estimator $\hat{\theta}$ is measured by $\Phi(\rho(\theta, \hat{\theta})) = \|\theta - \hat{\theta}\|^2$, where:
 - ▶ $\theta = \theta(P)$ is **expectation** of $P = N(\theta, I_d)$
 - ▶ $\hat{\theta} = \hat{\theta}(X_1^n)$ is the estimator of interest. Examples: $n^{-1}(\sum_{i=1}^n X_i)$, X_1 .
 - ▶ $\Phi(t) = t^2$ is a non-decreasing function
 - ▶ $\rho(\theta, \hat{\theta}) = \|\theta - \hat{\theta}\|$ is a semimetric
- ▶ Question: What would be the best performance of an ideal estimator in the worse case?

$$\mathcal{M}_n(\theta(\mathcal{P}), \Phi \circ \rho) := \inf_{\hat{\theta}} \sup_{P \in \mathcal{P}} \mathbb{E} \left[\Phi(\rho(\theta, \hat{\theta})) \right]$$

Finding exact $\mathcal{M}()$ is difficult, instead our attempt is to find a lower bound of it.

Recap: General Approach to Find Lower Bound

$$\mathcal{M}_n(\theta(\mathcal{P}), \Phi \circ \rho) := \inf_{\hat{\theta}} \sup_{P \in \mathcal{P}} \mathbb{E} \left[\Phi(\rho(\theta, \hat{\theta})) \right]$$

- ▶ Translate to probability (Markov inequality)

$$\inf_{\hat{\theta}} \sup_{P \in \mathcal{P}} \mathbb{E} \left[\Phi(\rho(\theta, \hat{\theta})) \right] \geq \Phi(\delta) \inf_{\hat{\theta}} \sup_{P \in \mathcal{P}} \mathbb{P}(\rho(\theta, \hat{\theta}) \geq \delta)$$

- ▶ Reduce the whole space \mathcal{P} to a finite set $\{\theta_v | v \in \mathcal{V}\}$

$$\inf_{\hat{\theta}} \sup_{P \in \mathcal{P}} \mathbb{P}(\rho(\theta, \hat{\theta}) \geq \delta) \geq \inf_{\hat{\theta}} \max_v \mathbb{P}(\rho(\theta_v, \hat{\theta}) \geq \delta)$$

- ▶ Reduce to a hypothesis testing error by constructing **2δ -packing set**.

$$\inf_{\hat{\theta}} \max_v \mathbb{P}(\rho(\theta_v, \hat{\theta}) \geq \delta) \geq \inf_{\Psi} \max_v \mathbb{P}(v \neq \Psi(\tilde{X}_1^n)),$$

where $\Psi(\tilde{X}_1^N) \triangleq \arg \min_v \rho(\theta_v, \hat{\theta}(\tilde{X}_1^N))$

Recap

- ▶ Fano's method is to switch to the average.

$$\begin{aligned}\inf_{\Psi} \max_v \mathbb{P}(v \neq \Psi(\tilde{X}_1^n)) &\geq \inf_{\Psi} \frac{1}{|\mathcal{V}|} \sum_v \mathbb{P}(v \neq \Psi(\tilde{X}_1^N)) \\ &= \inf_{\Psi} \mathbb{P}(V \neq \Psi(\tilde{X}_1^N)), \quad \text{where } V \text{ is a uniform RV.}\end{aligned}$$

Lemma

For any discrete RVs V, V' on the same alphabet \mathcal{V} ,

$$\mathbb{P}(V \neq V') \geq 1 - \frac{I(V; V') + \log 2}{\log |\mathcal{V}|}$$

where \mathbb{P} is taken with respect to both V, V' .

- ▶ There are other alternatives which do not consider RV V [Tsybakov 2009].

Recap: Fano's Method - The Recipe

We want to lower bound the RHS of

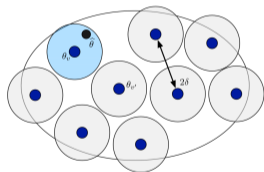
$$\mathbb{P}(V \neq V') \geq 1 - \frac{I(V; V') + \log 2}{\log |\mathcal{V}|}$$

by construct a **packing set** $\{\theta_v | v \in \mathcal{V}\}$, such that

- ▶ (required) $\rho(\theta_v, \theta_{v'}) \geq 2\delta \quad \forall v, v' \in \mathcal{V}$
- ▶ (desired) $|\mathcal{V}|$ is *large*
- ▶ (desired) $I(V; V')$ is *small*
- ▶ In the Gaussian mean estimation example, $|\mathcal{V}| \geq 2^d$, $I(V; V') \leq O(n\delta^2)$.

Two tasks:

- ▶ Construct packing set.
- ▶ Lower bound mutual information.



Minimax Bound in 1-bit Matrix Completion Problem

- ▶ Given matrix $\mathbf{M} \in K \triangleq \{\mathbf{M} \in \mathbb{R}^{d_1 \times d_2} \mid \|\mathbf{M}\|_* \leq \alpha\sqrt{rd_1d_2}, \|\mathbf{M}\|_\infty \leq \alpha\}$.
- ▶ A RV $\Omega \subset [d_1] \times [d_2]$ with $\mathbb{E}[|\Omega|] = n$
- ▶ A differential function $f : \mathbb{R} \rightarrow [1, 0]$ (cdf)
- ▶ Matrix \mathbf{Y} such that

$$Y_{ij} = \begin{cases} +1 & \text{with probability } f(M_{ij}) \\ -1 & \text{with probability } 1 - f(M_{ij}) \end{cases}$$

- ▶ Task: Estimate \mathbf{M} given \mathbf{Y}, Ω
- ▶ Quality measurement: $\Phi(\rho(\mathbf{M}, \widehat{\mathbf{M}})) = \frac{1}{d_1d_2} \|\mathbf{M} - \widehat{\mathbf{M}}\|_{\text{F}}^2$.

Theorem (Davenport et al. 2014)

Given a fixed algorithm, there exists $\mathbf{M} \in K$ such that with probability at least 0.75 (over RV \mathbf{Y}),

$$\frac{1}{d_1d_2} \|\mathbf{M} - \widehat{\mathbf{M}}\|_{\text{F}}^2 \geq \min \left(C_1, C_2\alpha\sqrt{\beta_{0.75\alpha}}\sqrt{\frac{r \max(d_1, d_2)}{n}} \right) = O \left(\frac{1}{\sqrt{n}} \right)$$

Prove by construction!

Sketch of Proof: Step 1

- ▶ Construct a set of matrices $\mathcal{X} = \{\mathbf{X}_v\}$ indexed by $v \in \mathcal{V}$ such that
 - ▶ $\mathcal{X} \subseteq K$
 - ▶ $\|\mathbf{X}_v - \mathbf{X}_{v'}\|_F^2 \geq \epsilon^2, \quad \forall v, v' \in \mathcal{V}$ for some $\epsilon > 0$.
- ▶ Uniformly choose a $V \in \mathcal{V}$, constructing $P(\cdot; \mathbf{X}_V)$, draw set X_1^N of N samples from that $P(\cdot; \mathbf{X}_V)$.
- ▶ Let ψ be the algorithm in the theorem: it uses data \mathbf{Y}, Ω and outputs $\widehat{\mathbf{M}}$.
- ▶ Let Ψ define as $\widehat{V} = \Psi((\mathbf{Y}, \Omega)) = \arg \min_{v \in \mathcal{V}} \rho(\widehat{\mathbf{M}}, \mathbf{X}_v)$.

By construction,

$$\mathbb{P}(V \neq \widehat{V}) = \mathbb{P}\left(\|\mathbf{X}_V - \mathbf{X}_{\widehat{V}}\|_F^2 \geq \epsilon^2\right)$$

Hence, the remaining part is to find a bound on the best possible prediction accuracy? i.e.,

$$\inf_{\Psi} \mathbb{P}(V \neq \widehat{V}) \geq q(\epsilon)$$

where \mathbb{P} is respect to RVs V, \mathbf{Y}_Ω .

Then we can claim that there exists \mathbf{M} , with probability at least $q(\epsilon)$,

$$\left\|\mathbf{M} - \widehat{\mathbf{M}}\right\|_F^2 \geq \epsilon^2$$

Sketch of Proof: Step 2

- ▶ Find the lower bound of $\inf_{\Psi} \mathbb{P}(V \neq \widehat{V})$
 - ▶ Fano's inequality: For any discrete RVs V, V' on the same alphabet \mathcal{V} ,

$$P(V \neq \widehat{V}) \geq 1 - \frac{I(V; \widehat{V}) + \log 2}{\log \mathcal{V}}$$

- ▶ $I(V; \widehat{V}) \leq I(V; \mathbf{Y}_{\Omega})$ since we have a Markov chain $V \rightarrow \mathbf{Y}_{\Omega} \rightarrow \widehat{V}$
- ▶ Bound the $I(V; \mathbf{Y}_{\Omega})$ [Scarlett et al. 2019]
 - ▶ Tensorization if all data points are i.i.d
 - ▶ Otherwise,

$$I(V; \mathbf{Y}_{\Omega}) \leq \max_{v, v'} D_{\text{kl}}(P(\cdot|v) || P(\cdot|v'))$$

- ▶ Upper bound that KL, which is application-dependent. Also, the ϵ should appear in this step.
- ▶ Integrating everything together, choosing ϵ to have a tight/meaningful bound.

Step 1: Construct an Attentive Packing set

Lemma

Let $\gamma \leq 1$ be such that $r/\gamma^2 \in \mathbb{N}$, and suppose that $r/\gamma^2 \leq d_1$. There is a set $\mathcal{X} \subset K$ with

$$|\mathcal{X}| \geq \exp\left(\frac{rd_2}{16\gamma^2}\right)$$

with the following properties:

- ▶ For all $\mathbf{X} \in \mathcal{X}$, each entry has $|X_{ij}| = \alpha\gamma$.
- ▶ For all $\mathbf{X} \neq \mathbf{X}' \in \mathcal{X}$,

$$\|\mathbf{X} - \mathbf{X}'\|_F^2 > 0.5\alpha^2\gamma^2d_1d_2$$

Proof of the Existence of Packing Set

It is an interesting probabilistic argument.

Consider the following distribution over random matrix \mathbf{X} with size of $d_1 \times d_2$:

- ▶ Let $d'_1 \triangleq r/\gamma^2 (\leq d_1)$.
- ▶ Matrix \mathbf{X} contains multiple blocks of size $d'_1 \times d_2$.
- ▶ For the first block, all entries are i.i.d Bernoulli RVs, i.e., $X_{ij} \sim \text{Bernoulli}(0.5)$, $X_{ij} \in \{\pm\alpha\gamma\}$, $\forall (i, j) \in [d'_1] \times [d_2]$.
- ▶ For other blocks are just copies of the first block (as much as possible).

We will draw from this distribution to construct set \mathcal{X} of $\left\lceil \exp\left(\frac{rd_2}{16\gamma^2}\right) \right\rceil$ elements.

Then $\mathcal{X} \subset K \triangleq \{\mathbf{M} \in \mathbb{R}^{d_1 \times d_2} \mid \|\mathbf{M}\|_* \leq \alpha\sqrt{rd_1d_2}, \|\mathbf{M}\|_\infty \leq \alpha\}$ since

- ▶ $\|\mathbf{X}\|_\infty = \alpha\gamma \leq \alpha$
- ▶ $\|\mathbf{X}\|_* \leq \sqrt{\text{rank}(\mathbf{X})} \|\mathbf{X}\|_F \leq \sqrt{d'_1} \|\mathbf{X}\|_F = \sqrt{r/\gamma^2} \sqrt{d_1d_2} \alpha\gamma = \alpha\sqrt{rd_1d_2}$

For 2 RVs \mathbf{X}, \mathbf{Y} followed the above distribution,

$$\begin{aligned}
 \|\mathbf{X} - \mathbf{Y}\|_{\text{F}}^2 &= \sum_{i,j} (X_{ij} - Y_{ij})^2 \\
 &\geq \left\lfloor \frac{d_1}{d'_1} \right\rfloor \sum_{i \in [d'_1], j \in [d_2]} (X_{ij} - Y_{ij})^2 \\
 &= 4\alpha^2\gamma^2 \left\lfloor \frac{d_1}{d'_1} \right\rfloor \sum_{i \in [d'_1], j \in [d_2]} \delta_{ij}, \quad \delta_{ij} \sim_{\text{i.i.d}} \text{Bern}(0.5), \delta_{ij} \in \{0, 1\}
 \end{aligned}$$

Next, with union bound and Hoeffding's inequality, we obtain,

$$\begin{aligned}
 P \left(\min_{\mathbf{X} \neq \mathbf{Y}} \sum_{i \in [d'_1], j \in [d_2]} \delta_{ij} \leq 0.25d'_1d_2 \right) &\leq \sum_{\mathbf{X} \neq \mathbf{Y}} P \left(\min_{\mathbf{X} \neq \mathbf{Y}} \sum_{i \in [d'_1], j \in [d_2]} \delta_{ij} \leq 0.25d'_1d_2 \right) \\
 &\leq \binom{|\mathcal{X}|}{2} \exp(-d'_1d_2/8) < 1
 \end{aligned}$$

That means that there is a non-zero probability that we obtain the set \mathcal{X} such that

$$\|\mathbf{X} - \mathbf{Y}\|_{\text{F}}^2 \geq \alpha^2\gamma^2 \left\lfloor \frac{d_1}{d'_1} \right\rfloor d'_1d_2 \geq 0.5\alpha^2\gamma^2d_1d_2$$

Step 2: Apply Fano's Inequality

Let $\mathcal{X}'_{\alpha/2,\gamma}$ be the set constructed in the previous Lemma. Construct \mathcal{X} as

$$\mathcal{X} \triangleq \left\{ \mathbf{X}' + \alpha \left(1 - \frac{\gamma}{2}\right) \mathbf{1} \mid \mathbf{X}' \in \mathcal{X}'_{\alpha/2,\gamma} \right\},$$

where γ is chosen as

$$4\sqrt{2}\epsilon/\alpha \leq \gamma \leq 8\epsilon/\alpha,$$

and ϵ is chosen such that

$$\|\mathbf{X} - \mathbf{X}'\|_{\text{F}}^2 \leq 4d_1d_2\epsilon^2$$

By construction, $\mathcal{X} \subset K$ (not obvious but easy to show), and $|\mathcal{X}| = |\mathcal{X}'_{\alpha/2,\gamma}|$.

Fano's Inequality

Now we show that if we choose $M \in \mathcal{X}$ uniformly

$$P(V \neq \hat{V}) \geq 1 - \frac{\max_{v,v'} D_{\text{kl}}(P(\cdot|v) || P(\cdot|v')) + \log 2}{\log |\mathcal{V}|}$$

By property of KL divergence of product distributions,

$$\max_{v,v' \in \mathcal{V}} D_{\text{kl}}(\mathbf{Y}_\Omega|v || \mathbf{Y}_\Omega|v') = \max_{v,v' \in \mathcal{V}} \sum_{(i,j) \in \Omega} D_{\text{kl}}(Y_{ij}|v || Y_{ij}|v')$$

- ▶ All summands are D_{KL} between 2 Bernoulli RVs
- ▶ They are either 0, $D_{\text{kl}}(\alpha||\alpha')$, $D_{\text{kl}}(\alpha'||\alpha)$ (because of our construction of the packing set).

Lemma

For $x, y \in (0, 1)$, $X \sim \text{Bern}(x)$, $Y \sim \text{Bern}(y)$. Then

$$D_{\text{kl}}(x||y) \leq \frac{(x - y)^2}{y(1 - y)}$$

Using the above Lemma,

$$\begin{aligned}
 D_{\text{kl}}(Y_{ij}|v \parallel Y_{ij}|v') &\leq \frac{(f(\alpha) - f(\alpha'))^2}{f(\alpha')(1 - f(\alpha'))} \\
 &\leq \frac{(f'(\xi))^2(\alpha - \alpha')^2}{f(\alpha')(1 - f(\alpha'))} \quad \text{for some } \xi \in [\alpha', \alpha] \quad (\text{intermediate value theorem}) \\
 &\leq \frac{(\gamma\alpha)^2}{\beta_{\alpha'}} \quad (\text{since } \alpha' = (1 - \gamma)\alpha) \\
 &\leq \frac{64\epsilon^2}{\beta_{\alpha'}} \quad (\text{by assumption}) \\
 &\Rightarrow I(V; \hat{V}) \leq \frac{64n\epsilon^2}{\beta_{\alpha'}}
 \end{aligned}$$

Hence,

$$\begin{aligned}
 \inf_{\Psi} \mathbb{P}(\Psi(\mathbf{Y}_{\Omega}) \neq V) &\geq 1 - \frac{I(V; \hat{V}) + \log 2}{\log |\mathcal{X}|} \\
 &\geq 1 - 1024\epsilon^2 \left(\frac{64n\epsilon^2/\beta_{\alpha'} + 1}{\alpha^2 r d_2} \right)
 \end{aligned}$$

$$\inf_{\Psi} \mathbb{P}(\Psi(\mathbf{Y}_{\Omega}) \neq V) \geq 1 - 1024\epsilon^2 \left(\frac{64n\epsilon^2/\beta_{\alpha'} + 1}{\alpha^2 r d_2} \right)$$

Recall that

$$\left\| \mathbf{M} - \widehat{\mathbf{M}} \right\|_{\text{F}}^2 \geq 4d_1 d_2 \epsilon^2$$

Lastly, choose ϵ so that we get a meaningful bound. Choose

$$\epsilon^2 = \dots$$

then they can conclude that

$$\left\| \mathbf{M} - \widehat{\mathbf{M}} \right\|_{\text{F}}^2 \geq O(1/\sqrt{n})$$

with probability at least 0.75.

Some Comments

- ▶ The proof does not take into account RV Ω
- ▶ Proof of existence of packing set using probabilistic is a nice approach
- ▶ Data samples does not need to be independent
- ▶ Fano's inequality is a key step in the general minimax bound derivation